

# Are the order statistics ordered? (revisited)

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<sup>1</sup>Partially supported by Ministerio de Ciencia e Innovación of Spain under grant PID2019-108079GB-C22/AEI/10.13039/501100011033. > < << >> >>> >>>>

## References

The conference is based on the following references:

- ▶ Navarro J., Durante F., Fernández-Sánchez J. (2021) Connecting copula properties with reliability properties of coherent systems. *Applied Stochastic Models in Business and Industry* 37, 496–512.
- ▶ Navarro J., Torrado N., del Águila Y. (2018). Comparisons between largest order statistics from multiple-outlier models with dependence. *Methodology and Computing in Applied Probability* 20, 411–433.
- ▶ Navarro J., Rychlik T. and Shaked M. (2007). Are the order statistics ordered? A Survey of Recent Results. *Communications in Statistics Theory and Methods* 36 (7), 1273–1290.

# Outline

## Preliminary results

Stochastic orders

Order statistics

Distortion representations

## Ordering properties

Case I: IID

Case II: ID

Cases III & IV: IND & GEN

## Main references

## Preliminary results

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- ▶  $m_X(t) = E(X - t|X > t)$  and  $m_Y(t) = E(Y - t|Y > t)$  mean residual life functions (MRL).

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- ▶ Likelihood ratio order:  $X \leq_{LR} Y \Leftrightarrow f_Y/f_X$  increases.
- ▶ Relationships:

$$\begin{array}{ccccc} X \leq_{LR} Y & \Rightarrow & X \leq_{HR} Y & \Rightarrow & X \leq_{MRL} Y \\ & & \Downarrow & & \Downarrow \\ & & X \leq_{ST} Y & \Rightarrow & E(X) \leq E(Y) \end{array}$$

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- ▶  $X_{n-k+1:n}$  lifetime of a k-out-of-n system.

# Basic properties

- ▶ Survival function (IID case):

$$\bar{F}_{i:n}(t) = \Pr(X_{i:n} > t) = \sum_{j=0}^{i-1} \binom{n}{j} F^j(t) \bar{F}^{n-j}(t).$$

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- ▶ Hazard rate function (IID case):

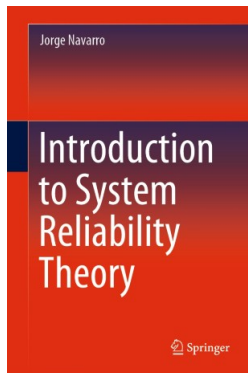
$$h_{i:n}(t) = i \binom{n}{i} \frac{f(t) F^{i-1}(t) \bar{F}^{n-i}(t)}{\sum_{j=0}^{i-1} \binom{n}{j} F^j(t) \bar{F}^{n-j}(t)}.$$

## Basic references on order statistics and systems

- ▶ Arnold B.C., Balakrishnan N., Nagaraja, H.N. *A First Course in Order Statistics*. SIAM, 2008.

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- ▶ My new book:



# Distortion representations

- ▶ As we have seen, in the IID case we have:

$$\bar{F}_{i:n}(t) = \bar{q}(\bar{F}(t))$$

where  $\bar{q} : [0, 1] \rightarrow [0, 1]$  is a distortion function, i.e.,  $\bar{q}$  is continuous, is increasing and satisfies  $\bar{q}(0) = 0$  and  $\bar{q}(1) = 1$ .

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- ▶ In the general case, from any  $(X_1, \dots, X_n)$  it can be written as

$$\bar{F}_{i:n}(t) = \bar{Q}(\bar{F}_1(t), \dots, \bar{F}_n(t)) \quad \text{for all } t \in \mathbb{R}, \quad (1.1)$$

where  $\bar{F}_i(t) = \Pr(X_i > t)$  and  $\bar{Q} : [0, 1]^n \rightarrow [0, 1]$  is a distortion function, i.e.,  $\bar{Q}$  is continuous, increasing and satisfies  $\bar{Q}(0, \dots, 0) = 0$  and  $\bar{Q}(1, \dots, 1) = 1$ .

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- ▶  $\bar{Q}$  only depends on  $i$ ,  $n$  and the survival copula  $\hat{C}$  obtained from Sklar's theorem to get:

$$\Pr(X_1 > x_1, \dots, X_n > x_n) = \hat{C}(\bar{F}_1(x_1), \dots, \bar{F}_n(x_n))$$

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- ▶ IID case:

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with  $\bar{q}(u) = u^n$  for  $u \in [0, 1]$ .

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- ▶ Navarro et al. ASMBI, 2013 and Navarro and Gomis ASMBI, 2016.

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- ▶  $T_1 \leq_{RHR} T_2$  for all  $\bar{F}_1, \dots, \bar{F}_r$  iff  $Q_2/Q_1$  is increasing in  $(0, 1)^n$ .

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$$F_1 \geq_{ST} \dots \geq_{ST} F_n$$

iff  $\bar{Q}_1 \leq \bar{Q}_2$  in  $D = \{(u_1, \dots, u_n) \in [0, 1]^n : u_1 \geq \dots \geq u_n\}$ .

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- ▶ If  $T_i$  has SF  $\bar{Q}_i(\bar{F}_1, \dots, \bar{F}_n)$ ,  $i = 1, 2$ , then:
- ▶  $T_1 \leq_{ST} T_2$  for all  $\bar{F}_1, \dots, \bar{F}_n$  such that

$$F_1 \geq_{ST} \dots \geq_{ST} F_n$$

iff  $\bar{Q}_1 \leq \bar{Q}_2$  in  $D = \{(u_1, \dots, u_n) \in [0, 1]^n : u_1 \geq \dots \geq u_n\}$ .

- ▶  $T_1 \leq_{HR} T_2$  for all  $\bar{F}_1, \dots, \bar{F}_n$  such that

$$F_1 \geq_{HR} \dots \geq_{HR} F_n \quad (1.3)$$

iff the function

$$H(v_1, \dots, v_n) = \frac{\bar{Q}_2(v_1, v_1 v_2, \dots, v_1 \dots v_n)}{\bar{Q}_1(v_1, v_1 v_2, \dots, v_1 \dots v_n)} \quad (1.4)$$

is decreasing in  $(0, 1)^n$ .

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- ▶  $T_1 \leq_{HR} T_2$  for all  $\bar{F}_1, \dots, \bar{F}_n$  such that

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- ▶ A similar result holds for the RHR order.

# Ordering properties

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- ▶ **Case III:**  $X_1, \dots, X_n$  are IND with SF  $\bar{F}_1, \dots, \bar{F}_n$ .
- ▶ **Case IV:**  $X_1, \dots, X_n$  are arbitrary (GENeral case) with SF  $\bar{F}_1, \dots, \bar{F}_n$  and SC  $\hat{C}$ .

## Case I: IID, ST order

- ▶ Are  $X_{1:n} \leq_{ST} \cdots \leq_{ST} X_{n:n}$  ordered?

## Case I: IID, ST order

- ▶ Are  $X_{1:n} \leq_{ST} \cdots \leq_{ST} X_{n:n}$  ordered?
- ▶ Yes, since

$$\bar{F}_{i:n}(t) = \sum_{j=0}^{i-1} \binom{n}{j} F^j(t) \bar{F}^{n-j}(t) \leq \sum_{j=0}^i \binom{n}{j} F^j(t) \bar{F}^{n-j}(t) = \bar{F}_{i+1:n}(t)$$

for  $i = 1, \dots, n-1$  and all  $n, F$ .

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- ▶ Are  $X_{1:n} \leq_{HR} \cdots \leq_{HR} X_{n:n}$  ordered?
- ▶ Yes, since

$$\frac{\bar{F}_{i+1:n}(t)}{\bar{F}_{i:n}(t)} = \frac{\sum_{j=0}^i \binom{n}{j} F^j(t) \bar{F}^{n-j}(t)}{\sum_{j=0}^{i-1} \binom{n}{j} F^j(t) \bar{F}^{n-j}(t)} = 1 + \frac{\binom{n}{i} F^i(t) \bar{F}^{n-i}(t)}{\sum_{j=0}^{i-1} \binom{n}{j} F^j(t) \bar{F}^{n-j}(t)}$$

is increasing in  $t$  for  $i = 1, \dots, n-1$  and all  $n, F$  because

$$\sum_{j=0}^{i-1} \binom{n}{j} F^{j-i}(t) \bar{F}^{i-j}(t) = \sum_{j=0}^{i-1} \binom{n}{j} \bar{H}^{i-j}(t)$$

where  $H(t) = \bar{F}(t)/F(t) = -1 + 1/F(t)$  is decreasing.



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## Case I: IID, LR order

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- ▶ Yes, since

$$\frac{\bar{f}_{i+1:n}(t)}{\bar{f}_{i:n}(t)} = \frac{(i+1) \binom{n}{i+1} f(t) F^i(t) \bar{F}^{n-i-1}(t)}{i \binom{n}{i} f(t) F^{i-1}(t) \bar{F}^{n-i}(t)} = \frac{c}{H(t)}$$

is increasing in  $t$  for  $i = 1, \dots, n-1$  and all  $n, F$  (because  $H$  is decreasing).

## Case II: ID, ST order

- ▶ Are  $X_{1:n} \leq_{ST} \cdots \leq_{ST} X_{n:n}$  ordered?

## Case II: ID, ST order

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- ▶ Yes, since

$$\bar{F}_{i:n}(t) = \bar{q}_{i:n}(\bar{F}(t)) \leq \bar{F}_{i+1:n}(t) = \bar{q}_{i+1:n}(\bar{F}(t))$$

and  $\bar{q}_{i:n} \leq \bar{q}_{i+1:n}$  for all copula  $\hat{C}$ .

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and  $\bar{q}_{i:n} \leq \bar{q}_{i+1:n}$  for all copula  $\hat{C}$ .

- ▶ For example, for  $n = 2$

$$\bar{F}_{1:2}(t) = \bar{q}_{1:2}(\bar{F}(t)) \leq \bar{F}_{2:2}(t) = \bar{q}_{2:2}(\bar{F}(t))$$

where  $\bar{q}_{1:2}(u) = \hat{C}(u, u)$ , and

$$\bar{F}_{2:2}(t) = \Pr(\max(X_1, X_2) > t) = \Pr(X_1 > t) + \Pr(X_2 > t) - \Pr(X_{1:2} > t)$$

that is,  $\bar{q}_{2:2}(u) = 2u - \hat{C}(u, u)$  with

$$\hat{C}(u, u) \leq 2u - \hat{C}(u, u)$$

since  $\hat{C}(u, u) \leq \hat{C}(1, u) = u$  for all copula  $\hat{C}$ .

## Case II: ID, HR order

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## Case II: ID, HR order

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- ▶ Are  $X_{1:2} \leq_{HR} X_{2:2}$  ordered?

## Case II: ID, HR order

- ▶ Are  $X_{1:n} \leq_{HR} \dots \leq_{HR} X_{n:n}$  ordered?
- ▶ Are  $X_{1:2} \leq_{HR} X_{2:2}$  ordered?
- ▶ It holds if and only if

$$\frac{\bar{q}_{2:2}(u)}{\bar{q}_{1:2}(u)} = \frac{2u - \hat{C}(u, u)}{\hat{C}(u, u)} = -1 + \frac{2u}{\hat{C}(u, u)}$$

is decreasing, that is, if and only if

$$r(u) = \frac{\hat{C}(u, u)}{u} = \frac{\delta_{\hat{C}}(u)}{u}$$

is increasing in  $(0, 1)$ .



## Case II: ID, HR order

- ▶ Are  $X_{1:n} \leq_{HR} \cdots \leq_{HR} X_{n:n}$  ordered?
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- ▶ Is this property true for any copula?

## Case II: ID, HR order

- ▶ Yes for the product copula

$$r(u) = \frac{\widehat{C}(u, u)}{u} = \frac{u^2}{u} = u.$$

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$$\widehat{C}(u, v) = \frac{uv}{u + v - uv}$$

then

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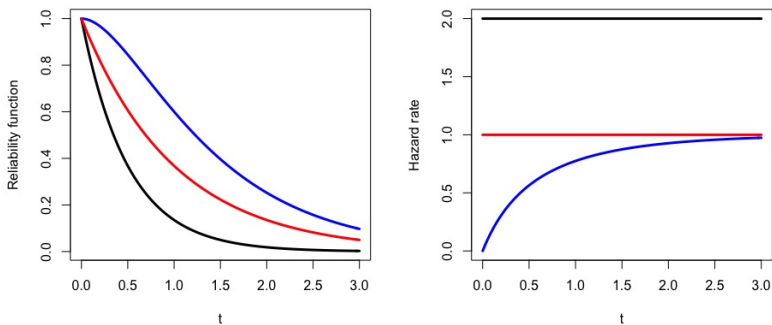
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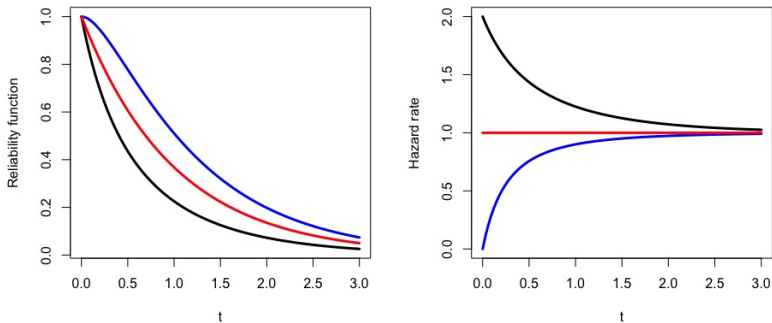
$$r(u) = \frac{C(u, u)}{u} = \frac{1}{2 - u}$$

which is increasing.

- ▶ Hence  $X_{1:2} \leq_{HR} X_{2:2}$  are ordered for any  $F$  and this copula.



**Figure:** Reliability and hazard rate functions of  $X_{1:2}$  (black line),  $X_i$  (red line) and  $X_{2:2}$  (blue line) when  $X_1$  and  $X_2$  are IID and have a common exponential distribution with mean one.



**Figure:** Reliability and hazard rate functions of  $X_{1:2}$  (black line),  $X_i$  (red line) and  $X_{2:2}$  (blue line) when  $X_1$  and  $X_2$  have a common exponential distribution with mean one and the Clayton copula given above.

## Case II: ID, HR order

- ▶ If we choose the following copula extracted from Example 4.1 in Navarro, Torrado and del Águila (2018)

$$\widehat{C}(u, v) = \min(u, v, 0.5\delta(u) + 0.5\delta(v))$$

with

$$\delta(u) = \begin{cases} u & \text{for } 0 \leq u \leq 1/3 \\ 1/3 & \text{for } 1/3 \leq u \leq 2/3 \\ 2u - 1 & \text{for } 2/3 \leq u \leq 1 \end{cases}$$

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- ▶ So the correct answer is NO (it depends on the copula).

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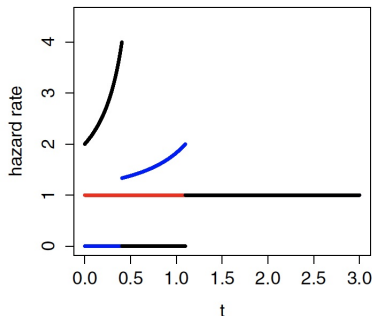
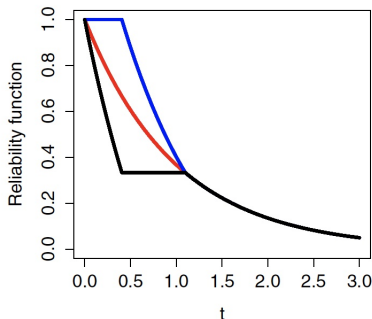
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- ▶ Then  $\delta_{\widehat{C}}(u) = \delta(u)$  and  $r(u) = \delta(u)/u$  is not increasing.
- ▶ So the correct answer is NO (it depends on the copula).
- ▶ This surprising fact was proved first in Navarro and Shaked, Journal of Applied Probability 43 (2006), 391-408.



**Figure:** Reliability and hazard rate functions of  $X_{1:2}$  (black line),  $X_i$  (red line) and  $X_{2:2}$  (blue line) when  $X_1$  and  $X_2$  have a common exponential distribution with mean one and the copula given above.

## Case II: ID, HR order

## Proposition (Navarro, Torrado and del Águila (2018))

Let  $X_1$  and  $X_2$  be the lifetimes of two components having a common distribution function  $F$  and copula and survival copula  $C$  and  $\widehat{C}$ , respectively. Then the following properties are equivalent:

- (i)  $X_{1:2} \leq_{HR} X_1$  for all  $F$ ;
- (ii)  $X_1 \leq_{HR} X_{2:2}$  for all  $F$ ;
- (iii)  $X_{1:2} \leq_{HR} X_{2:2}$  for all  $F$ ;
- (iv)  $\widehat{C}(u, u)/u$  is increasing in  $(0, 1)$ ;
- (v)  $(1 - C(u, u))/(1 - u)$  is increasing in  $(0, 1)$ .

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- ▶ Are  $X_{1:2} \leq_{LR} X_{2:2}$  ordered?
- ▶ It holds iff

$$\frac{\bar{q}'_{2:2}(u)}{\bar{q}'_{1:2}(u)} = \frac{2 - \delta'_{\hat{C}}(u)}{\delta'_{\hat{C}}(u)}$$

is decreasing in  $(0, 1)$ .

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- (iv)  $\widehat{C}(u, u)$  is convex in  $(0, 1)$ ;
- (v)  $C(u, u)$  is convex in  $(0, 1)$ .

## Cases III & IV: ST order, GEN case

- ▶ Are  $X_{1:n} \leq_{ST} \cdots \leq_{ST} X_{n:n}$  ordered in the GENERAL case?

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$$X_{1:n} \leq \cdots \leq X_{n:n}$$

implies

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see Shaked and Shanthikumar (2007), p. 5.

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see Shaked and Shanthikumar (2007), p. 5.

- ▶  $X_{i:m} \leq_{ST} X_{j:n}$  holds iff  $i \leq j$  and  $m - i \leq n - j$ , Arcones, Kvam and Samaniego, JASA, 2002

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- ▶  $X_{1:2} \leq_{HR} X_{2:2}$  holds iff

$$\frac{\bar{Q}_{2:2}(u, v)}{\bar{Q}_{1:2}(u, v)} = \frac{u + v - \hat{C}(u, v)}{\hat{C}(u, v)} \text{ is decreasing in } (0, 1)^2$$

that is, iff  $\hat{C}(u, v)/(u + v)$  is increasing.



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that is, iff  $\hat{C}(u, v)/(u + v)$  is increasing.

- ▶ If  $X_1$  and  $X_2$  are IND, then it holds iff

$$\frac{\bar{Q}_{2:2}(u, v)}{\bar{Q}_{1:2}(u, v)} = \frac{u + v - uv}{uv} = \frac{1}{u} + \frac{1}{v} - 1 \text{ is decreasing in } (0, 1)^2.$$

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- ▶ So the correct answer is YES (Boland, El-Newehi and Proschan, JAP, 1994).

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## Cases III & IV: HR order, IND case

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## Cases III & IV: HR order, IND case

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## Cases III &amp; IV: HR order, IND case

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- ▶  $X_{1:2} \leq_{HR} X_1$  holds iff

$$\frac{\bar{Q}_1(u, v)}{\bar{Q}_{1:2}(u, v)} = \frac{u}{\hat{C}(u, v)} \text{ is decreasing in } (0, 1)^2$$

that is, iff  $\hat{C}(u, v)/u$  is increasing in  $u$ , Navarro and Durante (2021).

## Cases III &amp; IV: HR order, IND case

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that is, iff  $\hat{C}(u, v)/u$  is increasing in  $u$ , Navarro and Durante (2021).

- ▶ If  $X_1$  and  $X_2$  are IND, then it holds since  $\hat{C}(u, v)/u = v$  is increasing (actually  $h_{1:2} = h_1 + h_2$ ).

## Cases III &amp; IV: HR order, IND case

- ▶ Are  $X_{1:2} \leq_{HR} X_1$  ordered in the GENERAL case?
- ▶ NO (proved before).
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- ▶ So the correct answer is YES.



## Cases III & IV: HR order, GEN case

### Proposition (Navarro and Durante (2021))

Let  $X_1$  and  $X_2$  be the lifetimes of two components having a distribution functions  $F_1$  and  $F_2$  and survival copula  $\widehat{C}$ . Then the following properties are equivalent:

- (i)  $X_{1:2} \leq_{HR} X_1$  for all  $F$ ;
- (ii)  $\widehat{C}(u, v)/u$  is increasing in  $u \in (0, 1)$  for all  $v \in (0, 1)$ ;
- (iii)  $(X_1, X_2)$  is Right Tail Decreasing  $RTD(X_2|X_1)$ , i.e.  $(X_2|X_1 > t)$  is ST decreasing in  $t$  (a negative dependence property).

## Cases III & IV: HR order, IND case

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$$\frac{\bar{Q}_{2:2}(u, v)}{\bar{Q}_1(u, v)} = \frac{u + v - uv}{u} = 1 + \frac{v}{u} - v \text{ is decreasing in } (0, 1)^2.$$

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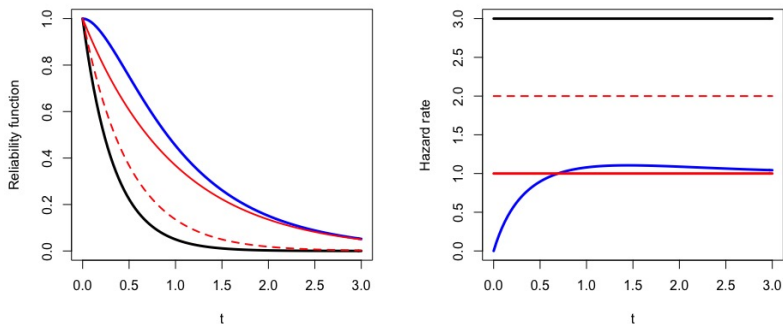
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- ▶ So the correct answer is NO ( $v(-1 + 1/u)$  is increasing in  $v$ ).



**Figure:** Reliability and hazard rate functions of  $X_{1:2}$  (black line),  $X_i$  (red lines) and  $X_{2:2}$  (blue line) when  $X_1$  and  $X_2$  are IND and have exponential distributions with mean one and two.



## Cases III & IV: HR order, GEN case

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Let  $X_1$  and  $X_2$  be the lifetimes of two components having a distribution functions  $F_1$  and  $F_2$  and survival copula  $\widehat{C}$ . Then the following properties are equivalent:

- (i)  $X_{1:2} \leq_{HR} X_{2:2}$  for all  $F$ ;
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## Cases III & IV: HR order, GEN case

- ▶ These properties hold if  $(X_1, X_2)$  is Right Tail Decreasing in both variables, i.e.,  $RTD(X_2|X_1)$  and  $RTD(X_1|X_2)$ .

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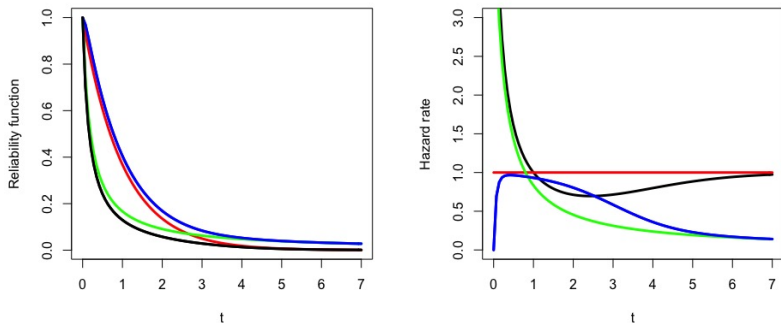
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- ▶ See Navarro and Durante (2021).



**Figure:** Reliability and hazard rate functions of  $X_{1:2}$  (black line),  $X_1$  (red line),  $X_3$  (green line) and  $X_{2:2}$  (blue line) when  $X_1$  and  $X_2$  are dependent with a survival Clayton copula and  $X_1$  has an exponential distribution with mean one and  $X_2$  has a Pareto distribution.

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- ▶ So the correct answer is NO.

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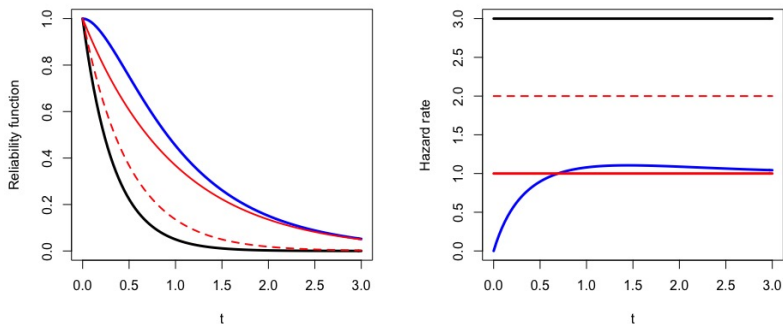
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## Main references

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- ▶ Questions?